

# M2 Complex Systems - Complex Networks

## Lecture 3 - Network models

Erdős-Rényi random graphs and configuration model

Autumn 2021 - ENS Lyon

Christophe Crespelle

`christophe.crespelle@ens-lyon.fr`

\* Thanks to Daron Acemoglu and Asu Ozdaglar for pedagogical material used for these slides.

# Network models

Model = **random generation** of synthetic networks

- To simulate :

- ▶ phenomena
- ▶ algorithms
- ▶ protocols

- In order to :

- ▶ design
- ▶ test
- ▶ predict
- ▶ better understand 

- Example :

Would Internet protocols still work if Internet was 10 times larger ?

- ▶ generate a synthetic network and simulate

# Network models

Model = **random generation** of synthetic networks  
... having the properties of real-world networks!!!

# Network models

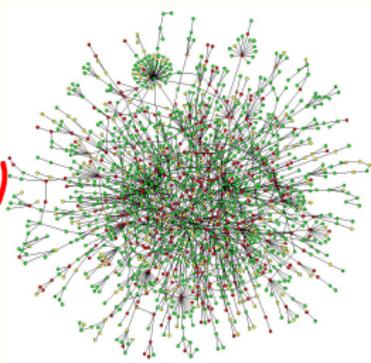
Model = **random generation** of synthetic networks  
... having the properties of real-world networks!!!

Four classic properties of real-world complex networks :

- Low global density

$$m = \# \text{ edges} \ll \frac{n(n-1)}{2}$$
$$\frac{m}{\frac{n(n-1)}{2}} \ll 1$$

$10^{-6}$     $10^{-5}$



# Network models

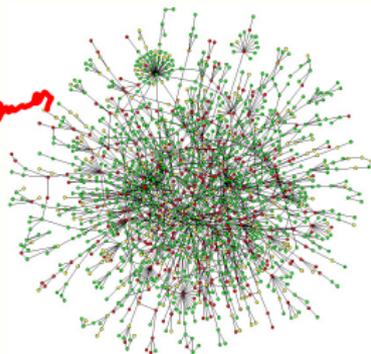
Model = **random generation** of synthetic networks  
... having the properties of real-world networks!!!

Four classic properties of real-world complex networks :

- Low global density
- Short distances

average distance  $\sim \log n$   
(diameter)  $\sim \log n$

$\rightarrow$  random

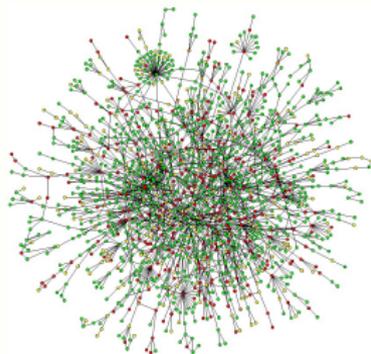


# Network models

Model = **random generation** of synthetic networks  
... having the properties of real-world networks!!!

Four classic properties of real-world complex networks :

- Low global density
- Short distances
- Heterogeneous degrees

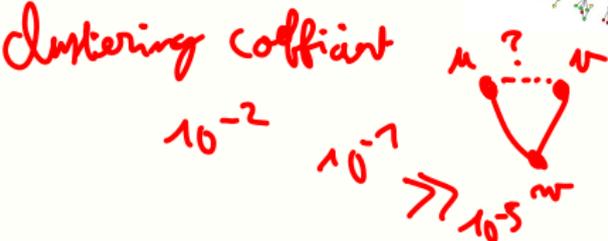
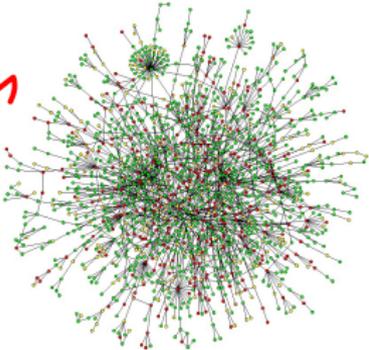


# Network models

Model = **random generation** of synthetic networks  
... having the properties of real-world networks!!!

Four classic properties of real-world complex networks :

- Low global density
  - Short distances
  - Heterogeneous degrees
  - High local density
- randomness*
- $\neq$



# Network models

Model = **random generation** of synthetic networks  
... having the properties of real-world networks!!!

Four classic properties of real-world complex networks :

*in notation.*

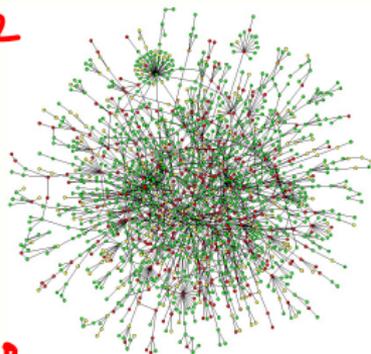
- Low global density
- Short distances
- Heterogeneous degrees
- High local density

*0,000342*

*3,45*



*clust coef = 0,08*



Goal : generate synthetic networks having these four properties  
(in a generic way)

# Erdős-Rényi (ER) random graphs

There are two models :

- $G_{n,m}$  : choose uniformly at random (u.a.r.)  $m$  edges among the  $n$  vertices

# Erdős-Rényi (ER) random graphs

There are two models :

- $G_{n,m}$  : choose uniformly at random (u.a.r.)  $m$  edges among the  $n$  vertices
- $G_{n,p}$  : for each couple of the  $n$  vertices, put an edge with probability  $p$

$$p \frac{n(n-1)}{2} = m$$

$$p = \frac{2m}{n(n-1)}$$

# Erdős-Rényi (ER) random graphs

There are two models :

- $G_{n,m}$  : choose uniformly at random (u.a.r.)  $m$  edges among the  $n$  vertices
- $G_{n,p}$  : for each couple of the  $n$  vertices, put an edge with probability  $p$

⇒ "essentially" equivalent when  $p = \frac{2m}{n(n-1)}$

# Erdős-Rényi (ER) random graphs

1959  $G_{n,m}$

There are two models :

- $G_{n,m}$  : choose uniformly at random (u.a.r.)  $m$  edges among the  $n$  vertices
- $G_{n,p}$  : for each couple of the  $n$  vertices, put an edge with probability  $p$

$\Rightarrow$  "essentially" equivalent when  $p = \frac{2m}{n(n-1)}$

Should we use  $G_{n,m}$  or  $G_{n,p}$ ?  $\rightarrow O(n^2)$

- For generating networks?  $G_{n,m}$   $O(m)$
- For mathematical analysis of the model?  $G_{n,p}$

# $G_{n,m}$ : implementation and complexity $O(m)$

- Algo : Pick  $m$  times two vertices uniformly at random

$0 \dots m-1$   $u = \text{Rand}(0, m-1)$   $(a, b)$   
 $(a, b)$   $b = \text{Rand}(0, m-1)$   
risk of collision.  $\rightarrow$  multiple edges  
 $(a, a)$  self loop

Pick until you get  $m$  distinct edges.

$O(m)$  time.

## $G_{n,m}$ : implementation and complexity

- Algo : Pick  $m$  times two vertices uniformly at random
  - ▶ How to deal with self-loops?

## $G_{n,m}$ : implementation and complexity

- Algo : Pick  $m$  times two vertices uniformly at random
  - ▶ How to deal with self-loops ?
  - ▶ How to deal with multiple edges ?

# Properties of $G_{n,p}$

Four properties to check :

- Low global density

▶  $p$  parameter of the model, controls  $m$  :  $\mathbb{E}(m) = \frac{pn(n-1)}{2}$

# Properties of $G_{n,p}$

Four properties to check :

- Low global density ✓
  - ▶  $p$  parameter of the model, controls  $m$  :  $\mathbb{E}(m) = \frac{pn(n-1)}{2}$
  - ▶ law of large numbers :  $m$  is very concentrated around its mean

## Properties of $G_{n,p}$

Four properties to check :

- Low global density ✓
- Short distances

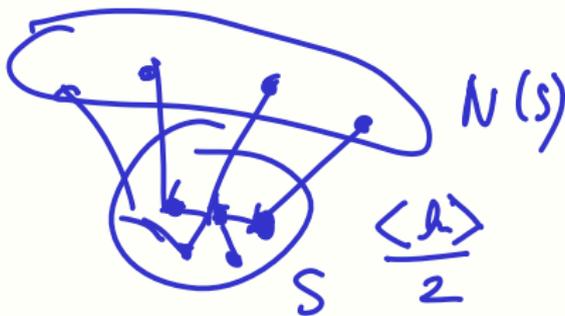
# Properties of $G_{n,p}$

Four properties to check :

- Low global density ✓
- Short distances

## Expansion property

- ▶ def. (graph theory) : a graph  $G$  is a  $c$ -vertex-expander iff
- $\forall S \subseteq V$  s.t.  $|S| \leq \frac{|V(G)|}{2}$ , we have  $|N(S)| \geq \underbrace{c \cdot |S|}$



# Properties of $G_{n,p}$

Four properties to check :

- Low global density ✓
- Short distances

## Expansion property

- ▶ def. (graph theory) : a graph  $G$  is a  $c$ -vertex-expander iff  $\forall S \subseteq V$  s.t.  $|S| \leq \frac{|V(G)|}{2}$ , we have  $|N(S)| \geq c \cdot |S|$
- ▶ expansion of  $G_{n,m}$  ?

# Properties of $G_{n,p}$

Four properties to check :

- Low global density ✓
- Short distances

## Expansion property

- ▶ def. (graph theory) : a graph  $G$  is a  $c$ -vertex-expander iff  $\forall S \subseteq V$  s.t.  $|S| \leq \frac{|V(G)|}{2}$ , we have  $|N(S)| \geq c \cdot |S|$
- ▶ expansion of  $G_{n,m}$ ?  $\sim \frac{m}{n}$   $\frac{2m}{n} = \langle \ell \rangle$

# Properties of $G_{n,p}$

Four properties to check :

- Low global density ✓
- Short distances ✓

## Expansion property

- ▶ def. (graph theory) : a graph  $G$  is a  $c$ -vertex-expander iff  $\forall S \subseteq V$  s.t.  $|S| \leq \frac{|V(G)|}{2}$ , we have  $|N(S)| \geq c \cdot |S|$
- ▶ expansion of  $G_{n,m}$ ?  $\sim \frac{m}{n}$
- ▶ until  $\frac{n}{2}$ , exponential growth of  $|B(u, d)| \sim \underbrace{(1+c)^d}$

# Properties of $G_{n,p}$

Four properties to check :

- Low global density ✓
- Short distances ✓

## Expansion property

- ▶ def. (graph theory) : a graph  $G$  is a  $c$ -vertex-expander iff  $\forall S \subseteq V$  s.t.  $|S| \leq \frac{|V(G)|}{2}$ , we have  $|N(S)| \geq c \cdot |S|$
  - ▶ expansion of  $G_{n,m}$ ?  $\sim \frac{m}{n}$
  - ▶ until  $\frac{n}{2}$ , exponential growth of  $|B(u, d)| \sim (1 + c)^d$
- Heterogeneous degrees

# Properties of $G_{n,p}$

$p(m)$

Four properties to check :

- Low global density ✓
- Short distances ✓

## Expansion property

- ▶ def. (graph theory) : a graph  $G$  is a  $c$ -vertex-expander iff  $\forall S \subseteq V$  s.t.  $|S| \leq \frac{|V(G)|}{2}$ , we have  $|N(S)| \geq c \cdot |S|$
- ▶ expansion of  $G_{n,m}$ ?  $\sim \frac{m}{n}$
- ▶ until  $\frac{n}{2}$ , exponential growth of  $|B(u, d)| \sim (1 + c)^d$
- Heterogeneous degrees
  - ▶ fix the average degree  $\lambda = p(n - 1)$

# Properties of $G_{n,p}$

Four properties to check :

- Low global density ✓

- Short distances ✓

## Expansion property

- ▶ def. (graph theory) : a graph  $G$  is a  $c$ -vertex-expander iff

$$\forall S \subseteq V \text{ s.t. } |S| \leq \frac{|V(G)|}{2}, \text{ we have } |N(S)| \geq c \cdot |S|$$

- ▶ expansion of  $G_{n,m}$ ?  $\sim \frac{m}{n}$

- ▶ until  $\frac{n}{2}$ , exponential growth of  $|B(u, d)| \sim (1+c)^d$

- Heterogeneous degrees

- ▶ fix the average degree  $\lambda = p(n-1)$

- ▶  $\mathbb{P}(d^o = k) = \binom{n-1}{k} p^k (1-p)^{n-1-k}$

$$= \frac{A_n^k}{k!} \frac{\lambda^k}{(n-1)^k} \left(1 - \frac{\lambda}{n-1}\right)^{n-1-k}$$

$$= \frac{A_n^k}{(n-1)^k} \frac{\lambda^k}{k!} \left(1 - \frac{\lambda}{n-1}\right)^{n-1} \left(1 - \frac{\lambda}{n-1}\right)^{-k}$$

$\downarrow$   
 $\lambda$

$\downarrow$   
 $e^{-\lambda}$

$\downarrow$   
 $\lambda$

$\lambda$ : fix  
 $n \rightarrow +\infty$

# Properties of $G_{n,p}$

normal

Four properties to check :

- Low global density ✓

- Short distances ✓

## Expansion property

- ▶ def. (graph theory) : a graph  $G$  is a  $c$ -vertex-expander iff

$$\forall S \subseteq V \text{ s.t. } |S| \leq \frac{|V(G)|}{2}, \text{ we have } |N(S)| \geq c \cdot |S|$$

- ▶ expansion of  $G_{n,m}$ ?  $\sim \frac{m}{n}$

- ▶ until  $\frac{n}{2}$ , exponential growth of  $|B(u, d)| \sim (1+c)^d$

- Heterogeneous degrees ✗

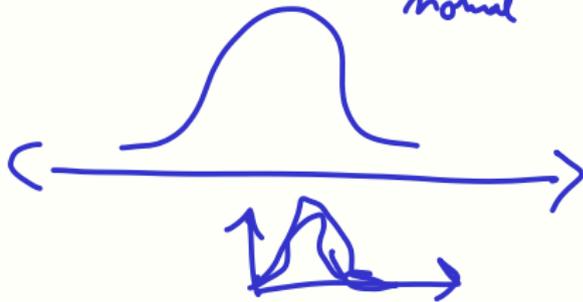
- ▶ fix the average degree  $\lambda = p(n-1)$

- ▶  $\mathbb{P}(d^\circ = k) = \binom{n-1}{k} p^k (1-p)^{(n-1-k)}$

$$= \frac{A_n^k}{k!} \frac{\lambda^k}{(n-1)^k} \left(1 - \frac{\lambda}{n-1}\right)^{n-1-k}$$

$$= \frac{A_n^k}{(n-1)^k} \frac{\lambda^k}{k!} \left(1 - \frac{\lambda}{n-1}\right)^{n-1} \left(1 - \frac{\lambda}{n-1}\right)^{-k}$$

- ▶ then when  $n \rightarrow +\infty$ ,  $\mathbb{P}(d^\circ = k) \rightarrow \frac{\lambda^k e^{-\lambda}}{k!}$ : Poisson law



# Properties of $G_{n,p}$

Four properties to check :

- Low global density ✓

- Short distances ✓

## Expansion property

- ▶ def. (graph theory) : a graph  $G$  is a  $c$ -vertex-expander iff

$\forall S \subseteq V$  s.t.  $|S| \leq \frac{|V(G)|}{2}$ , we have  $|N(S)| \geq c \cdot |S|$

- ▶ expansion of  $G_{n,m}$ ?  $\sim \frac{m}{n}$

- ▶ until  $\frac{n}{2}$ , exponential growth of  $|B(u, d)| \sim (1 + c)^d$

- Heterogeneous degrees ✗

- High local density

- ▶ probability of an edge in the neighbourhood of a vertex?



# Properties of $G_{n,p}$

Four properties to check :

- Low global density ✓
- Short distances ✓

## Expansion property

- ▶ def. (graph theory) : a graph  $G$  is a  $c$ -vertex-expander iff  $\forall S \subseteq V$  s.t.  $|S| \leq \frac{|V(G)|}{2}$ , we have  $|N(S)| \geq c \cdot |S|$
- ▶ expansion of  $G_{n,m}$ ?  $\sim \frac{m}{n}$
- ▶ until  $\frac{n}{2}$ , exponential growth of  $|B(u, d)| \sim (1 + c)^d$
- Heterogeneous degrees ✗
- High local density ✗
  - ▶ probability of an edge in the neighbourhood of a vertex?
  - ▶ same as everywhere :  $p$  (couples of vertices are independant)

# Phase transitions in $G_{n,p}$

N.B. :  $p$  (eventually) depends on  $n$

- Threshold function  $t(n)$  for property  $A$  :

▶  $\mathbb{P}(A) \rightarrow 0$  if  $\frac{p(n)}{t(n)} \rightarrow 0$

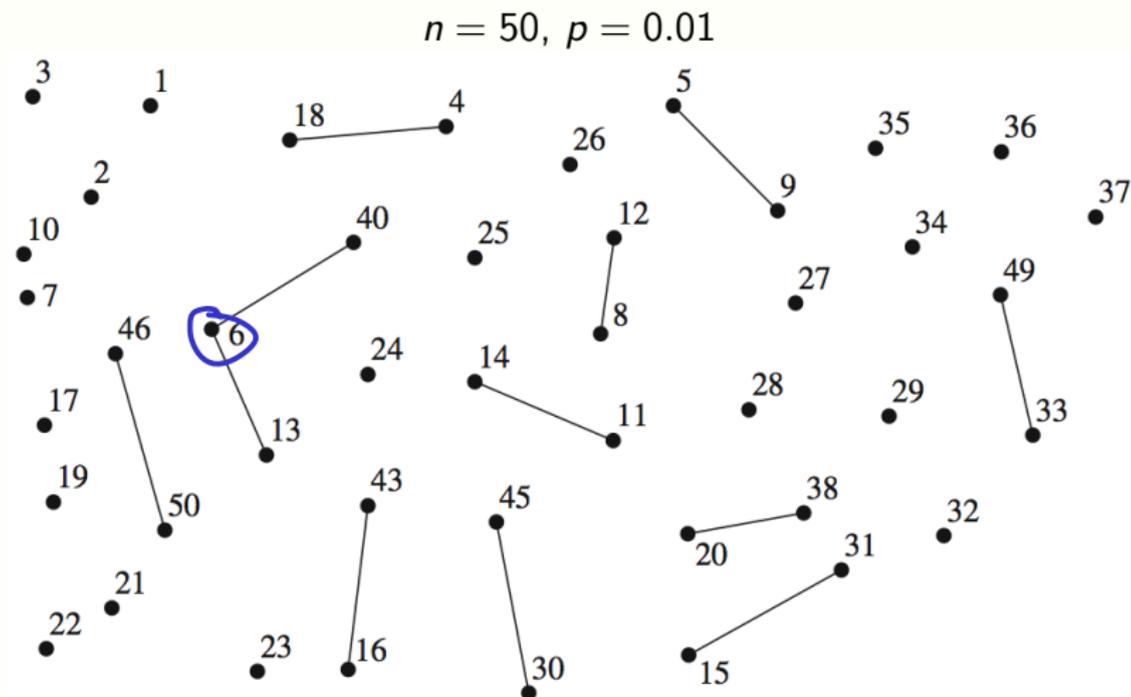
▶  $\mathbb{P}(A) \rightarrow 1$  if  $\frac{p(n)}{t(n)} \rightarrow +\infty$

- ▶ makes sense for monotonic properties (for inclusion of edge set)

$$p(n) = o(t(n))$$
$$t(n) = o(p(n))$$

- such a threshold function exists  $\Rightarrow$  phase transition
- Seminal work of Erdős and Rényi in 1959

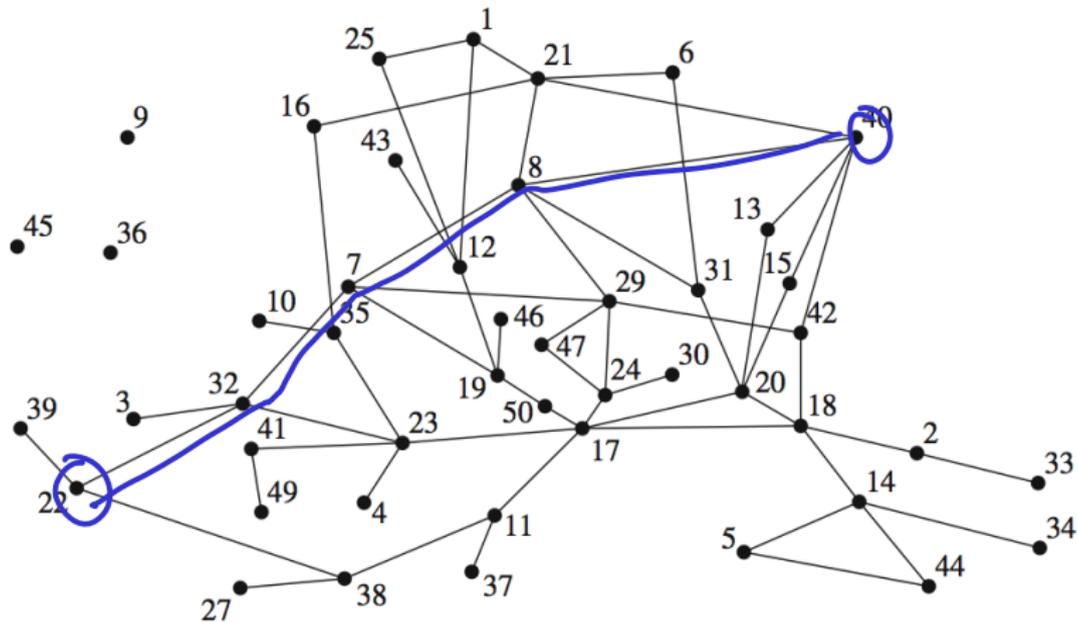
# Phase transitions in $G_{n,p}$





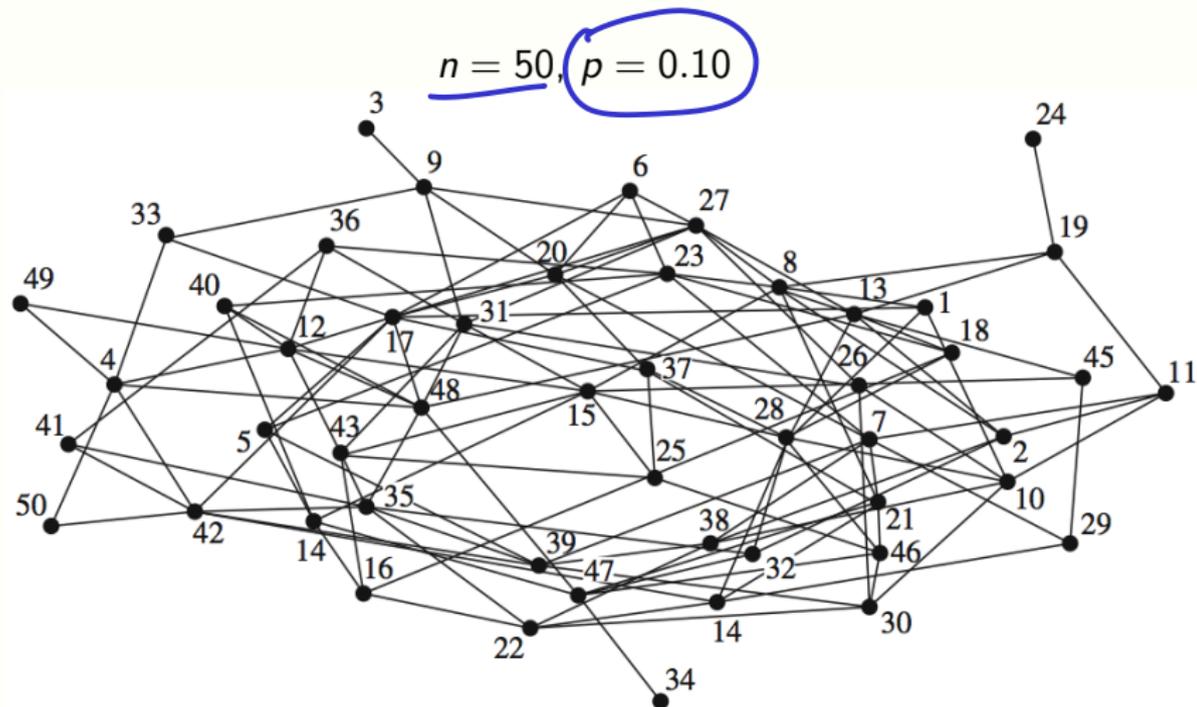
# Phase transitions in $G_{n,p}$

$n = 50, p = 0.05$



*single component.*

# Phase transitions in $G_{n,p}$



# Threshold for connectivity

- We show a threshold with function  $t(n) = \frac{\log n}{n}$
- Denote  $p(n) = \lambda \frac{\log n}{n}$  (mean degree  $\sim \lambda \log n$ )
- We show a (much) stronger statement for threshold function  $\frac{\log n}{n}$  :
  1.  $\mathbb{P}(\text{connectivity}) \rightarrow 0$  if  $\lambda < 1$
  2.  $\mathbb{P}(\text{connectivity}) \rightarrow 1$  if  $\lambda > 1$

## Proof of (1)

$$\lambda < 1$$

$$p(m) = \lambda \frac{\log m}{m}$$

- to prove (1), it is enough to show that the probability of existence of at least one isolated node goes to 1

0

## Proof of (1)

- to prove (1), it is enough to show that the probability of existence of at least one isolated node goes to 1
- Let  $I_i$  be the Bernoulli random variable defined as
  - ▶  $I_i = 1$  if vertex  $i$  is isolated
  - ▶  $I_i = 0$  otherwise

## Proof of (1)

- to prove (1), it is enough to show that the probability of existence of at least one isolated node goes to 1
- Let  $I_i$  be the Bernoulli random variable defined as
  - ▶  $I_i = 1$  if vertex  $i$  is isolated
  - ▶  $I_i = 0$  otherwise
- probability that a vertex is isolated :  
 $q = \mathbb{P}(I_i = 1) = (1 - p)^{n-1} \sim e^{-\lambda \log n} = n^{-\lambda}$

Handwritten blue annotations showing the asymptotic approximation of the probability of a vertex being isolated. The expression  $(1-p)^{n-1}$  is circled in blue, with a blue arrow pointing to the approximation  $e^{-\lambda \log n}$ . Below the circle, the letter  $n$  is written. To the right, the expression  $e^{-\lambda \log n}$  is written again.

## Proof of (1)

- to prove (1), it is enough to show that the probability of existence of at least one isolated node goes to 1
- Let  $I_i$  be the Bernoulli random variable defined as
  - ▶  $I_i = 1$  if vertex  $i$  is isolated
  - ▶  $I_i = 0$  otherwise
- probability that a vertex is isolated :  
 $q = \mathbb{P}(I_i = 1) = (1 - p)^{n-1} \sim e^{-\lambda \log n} = n^{-\lambda}$
- Let  $X = \sum_{i=1}^n I_i$  be the number of isolated vertices

# Proof of (1)

- to prove (1), it is enough to show that the probability of existence of at least one isolated node goes to 1
- Let  $I_i$  be the Bernoulli random variable defined as
  - ▶  $I_i = 1$  if vertex  $i$  is isolated
  - ▶  $I_i = 0$  otherwise
- probability that a vertex is isolated :  
 $q = \mathbb{P}(I_i = 1) = (1 - p)^{n-1} \sim e^{-\lambda \log n} = n^{-\lambda}$
- Let  $X = \sum_{i=1}^n I_i$  be the number of isolated vertices
- We have  $\mathbb{E}[X] = n \cdot n^{-\lambda} \rightarrow +\infty$  for  $\lambda < 1$

$$\frac{n}{n^\lambda}$$

## Proof of (1)

- to prove (1), it is enough to show that the probability of existence of at least one isolated node goes to 1
- Let  $I_i$  be the Bernoulli random variable defined as
  - ▶  $I_i = 1$  if vertex  $i$  is isolated
  - ▶  $I_i = 0$  otherwise
- probability that a vertex is isolated :  
 $q = \mathbb{P}(I_i = 1) = (1 - p)^{n-1} \sim e^{-\lambda \log n} = n^{-\lambda}$
- Let  $X = \sum_{i=1}^n I_i$  be the number of isolated vertices
- We have  $\mathbb{E}[X] = n \cdot n^{-\lambda} \rightarrow +\infty$  for  $\lambda < 1$
- enough to conclude that  $\mathbb{P}(X = 0) \rightarrow 0$ ?

## Proof of (1)

- to prove (1), it is enough to show that the probability of existence of at least one isolated node goes to 1
- Let  $I_i$  be the Bernoulli random variable defined as
  - ▶  $I_i = 1$  if vertex  $i$  is isolated
  - ▶  $I_i = 0$  otherwise
- probability that a vertex is isolated :  
 $q = \mathbb{P}(I_i = 1) = (1 - p)^{n-1} \sim e^{-\lambda \log n} = n^{-\lambda}$
- Let  $X = \sum_{i=1}^n I_i$  be the number of isolated vertices
- We have  $\mathbb{E}[X] = n \cdot n^{-\lambda} \rightarrow +\infty$  for  $\lambda < 1$
- enough to conclude that  $\mathbb{P}(X = 0) \rightarrow 0$  ?  
 $\Rightarrow$  NO, we need a concentration property.

## Proof of (1)

- We have  
 $\text{var}(X)$

$$X = \sum_i I_i$$

$$\begin{aligned} &= \sum_i \text{var}(I_i) + \sum_i \sum_{j \neq i} \text{cov}(I_i, I_j) \\ &= n \text{var}(I_1) + \underbrace{n(n-1) \text{cov}(I_1, I_2)} \end{aligned}$$

# Proof of (1)

$$E(X) = \frac{n}{2}$$

- We have
 
$$\begin{aligned} \text{var}(X) &= \sum_i \text{var}(I_i) + \sum_i \sum_{j \neq i} \text{cov}(I_i, I_j) \\ &= n \text{var}(I_1) + n(n-1) \text{cov}(I_1, I_2) \end{aligned}$$

$$E(X^2) \sim n^3$$

- And we also have :

$$\text{var}(I_1) = \mathbb{E}[I_1^2] - \mathbb{E}[I_1]^2 = q - q^2$$

$$\text{cov}(I_1, I_2) = \mathbb{E}[I_1 I_2] - \mathbb{E}[I_1] \mathbb{E}[I_2]$$

$$\mathbb{E}[I_1 I_2] = \mathbb{P}(I_1 = 1, I_2 = 1) = (1-p)^{2n-3} = \frac{q^2}{1-p}$$

$$\frac{n}{2} \uparrow$$

$$q = (1-p)^{n-1}$$

$$i^k \sim n^{k+1}$$

$n-2$

$2(n-2) + 1$

$$\sum_{i=1}^n i^2 \sim n^3$$

$$\sum_{i=1}^n i \sim n^2$$

## Proof of (1)

- We have
$$\begin{aligned} \text{var}(X) &= \sum_i \text{var}(I_i) + \sum_i \sum_{j \neq i} \text{cov}(I_i, I_j) \\ &= n \text{var}(I_1) + n(n-1) \text{cov}(I_1, I_2) \end{aligned}$$
- And we also have :
  - ▶  $\text{var}(I_1) = \mathbb{E}[I_1^2] - \mathbb{E}[I_1]^2 = q - q^2$
  - ▶  $\text{cov}(I_1, I_2) = \mathbb{E}[I_1 I_2] - \mathbb{E}[I_1] \mathbb{E}[I_2]$
  - ▶  $\mathbb{E}[I_1 I_2] = \mathbb{P}(I_1 = 1, I_2 = 1) = (1-p)^{2n-3} = \frac{q^2}{1-p}$
- We then obtain  $\text{var}(X) = nq(1-q) + n(n-1) \frac{q^2}{1-p}$

# Proof of (1)

$$p = \lambda \frac{\log n}{n}$$

- We have
$$\begin{aligned} \text{var}(X) &= \sum_i \text{var}(I_i) + \sum_i \sum_{j \neq i} \text{cov}(I_i, I_j) \\ &= n \text{var}(I_1) + n(n-1) \text{cov}(I_1, I_2) \end{aligned}$$

- And we also have :

- ▶  $\text{var}(I_1) = \mathbb{E}[I_1^2] - \mathbb{E}[I_1]^2 = q - q^2$

- ▶  $\text{cov}(I_1, I_2) = \mathbb{E}[I_1 I_2] - \mathbb{E}[I_1] \mathbb{E}[I_2]$

- ▶  $\mathbb{E}[I_1 I_2] = \mathbb{P}(I_1 = 1, I_2 = 1) = (1-p)^{2n-3} = \frac{q^2}{1-p}$

- We then obtain  $\text{var}(X) = nq(1-q) + n(n-1) \frac{q^2 p}{1-p}$

- when  $n \rightarrow +\infty$ , then  $q \rightarrow 0$  and  $p \rightarrow 0$

$$d = (1 - \lambda \frac{\log n}{n})^{n-1} \leq e^{-\lambda \log n}$$

## Proof of (1)

- We have
$$\begin{aligned} \text{var}(X) &= \sum_i \text{var}(l_i) + \sum_i \sum_{j \neq i} \text{cov}(l_i, l_j) \\ &= n \text{var}(l_1) + n(n-1) \text{cov}(l_1, l_2) \end{aligned}$$
- And we also have :
  - ▶  $\text{var}(l_1) = \mathbb{E}[l_1^2] - \mathbb{E}[l_1]^2 = q - q^2$
  - ▶  $\text{cov}(l_1, l_2) = \mathbb{E}[l_1 l_2] - \mathbb{E}[l_1] \mathbb{E}[l_2]$
  - ▶  $\mathbb{E}[l_1 l_2] = \mathbb{P}(l_1 = 1, l_2 = 1) = (1-p)^{2n-3} = \frac{q^2}{1-p}$
- We then obtain  $\text{var}(X) = nq(1-q) + n(n-1) \frac{q^2 p}{1-p}$
- when  $n \rightarrow +\infty$ , then  $q \rightarrow 0$  and  $p \rightarrow 0$
- this gives
$$\begin{aligned} \text{var}(X) &\sim nq + n^2 q^2 p \\ &= nn^{-\lambda} + \lambda n \log nn^{-2\lambda} \\ &\sim \underbrace{nn^{-\lambda}} = \underbrace{\mathbb{E}[X]} \end{aligned}$$

## Proof of (1)

- so we have  $\text{var}(X) \sim \mathbb{E}[X]$  |
- and because  $\text{var}(X) \geq (0 - \mathbb{E}[X])^2 \mathbb{P}(X = 0)$
- we obtain  $\mathbb{P}(X = 0) \leq \frac{1}{\mathbb{E}[X]} \rightarrow 0$
- it follows that  $\mathbb{P}(X > 0) \rightarrow 1$  when  $n \rightarrow +\infty$
- and consequently  $\mathbb{P}(\text{disconnected}) \rightarrow 1$  when  $n \rightarrow +\infty$

Proof of (2)  $\lambda < 1$



- we now fix  $\lambda > 1$
- let's check that  $\mathbb{E}[X] = nn^{-\lambda} \rightarrow 0$  when  $n \rightarrow +\infty$  ✓

## Proof of (2)

- we now fix  $\lambda > 1$
- let's check that  $\mathbb{E}[X] = nn^{-\lambda} \rightarrow 0$  when  $n \rightarrow +\infty$  ✓
- observe that  $G$  is disconnected  $\iff \exists k$  vertices without edges to the other vertices, for some  $k \leq n/2$

## Proof of (2)

- we now fix  $\lambda > 1$
- let's check that  $\mathbb{E}[X] = nn^{-\lambda} \rightarrow 0$  when  $n \rightarrow +\infty$  ✓
- observe that  $G$  is disconnected  $\iff \exists k$  vertices without edges to the other vertices, for some  $k \leq n/2$
- we have

$$\mathbb{P}(\{1, \dots, k\} \text{ not connected to the rest}) = (1 - p)^{k(n-k)}$$

## Proof of (2)

- we now fix  $\lambda > 1$
- let's check that  $\mathbb{E}[X] = nn^{-\lambda} \rightarrow 0$  when  $n \rightarrow +\infty$  ✓
- observe that  $G$  is disconnected  $\iff \exists k$  vertices without edges to the other vertices, for some  $k \leq n/2$
- we have
$$\mathbb{P}(\{1, \dots, k\} \text{ not connected to the rest}) = (1 - p)^{k(n-k)}$$
- and so
$$\mathbb{P}(\exists k \text{ vertices not connected to the rest}) \leq \binom{n}{k} (1 - p)^{k(n-k)}$$

## Proof of (2)

- we now fix  $\lambda > 1$
- let's check that  $\mathbb{E}[X] = nn^{-\lambda} \rightarrow 0$  when  $n \rightarrow +\infty$  ✓
- observe that  $G$  is disconnected  $\iff \exists k$  vertices without edges to the other vertices, for some  $k \leq n/2$
- we have  
$$\mathbb{P}(\{1, \dots, k\} \text{ not connected to the rest}) = (1 - p)^{k(n-k)}$$
- and so  $\cup$   
$$\mathbb{P}(\exists k \text{ vertices not connected to the rest}) \leq \binom{n}{k} (1 - p)^{k(n-k)}$$
- and finally  $\mathbb{P}(G \text{ is disconnected}) \leq \sum_{k=1}^{n/2} \binom{n}{k} (1 - p)^{k(n-k)}$

## Proof of (2)

- we now fix  $\lambda > 1$
- let's check that  $\mathbb{E}[X] = nn^{-\lambda} \rightarrow 0$  when  $n \rightarrow +\infty$  ✓
- observe that  $G$  is disconnected  $\iff \exists k$  vertices without edges to the other vertices, for some  $k \leq n/2$
- we have
$$\mathbb{P}(\{1, \dots, k\} \text{ not connected to the rest}) = (1 - p)^{k(n-k)}$$
- and so
$$\mathbb{P}(\exists k \text{ vertices not connected to the rest}) \leq \binom{n}{k} (1 - p)^{k(n-k)}$$
- and finally  $\mathbb{P}(G \text{ is disconnected}) \leq \sum_{k=1}^{n/2} \binom{n}{k} (1 - p)^{k(n-k)}$
- using this expression, one can show that
$$\mathbb{P}(G \text{ is disconnected}) \rightarrow 0 \text{ when } n \rightarrow +\infty$$

# Threshold for giant component

$$\langle h \rangle = \frac{2m}{n} \stackrel{\text{average degree}}{=} \rho_x(\text{critical})$$

$$P(h) \sim \frac{\log n}{n} \lambda^{(n-1)}$$

*average degree*

- Giant = constant fraction of the vertices *in a line  $\langle h \rangle \sim 2$*
- We show a threshold with function  $t(n) = \frac{1}{n}$   *$\lambda > 1$*
- Denote  $p(n) = \frac{\lambda}{n}$  (mean degree  $\sim \lambda$ )  *$\lambda < 1$*
- We again show a strong statement for threshold function  $\frac{1}{n}$  :
  1. if  $\lambda < 1$ ,  $\forall a \in \mathbb{R}_+^*$ ,  $\mathbb{P}(\text{maxsize}(CC) \geq a \log n) \rightarrow 0$
  2. if  $\lambda > 1$ ,  $\exists b \in \mathbb{R}_+^*$ ,  $\mathbb{P}(\text{maxsize}(CC) \geq b \cdot n) \rightarrow 1$

# Proof of (1) – preliminaries

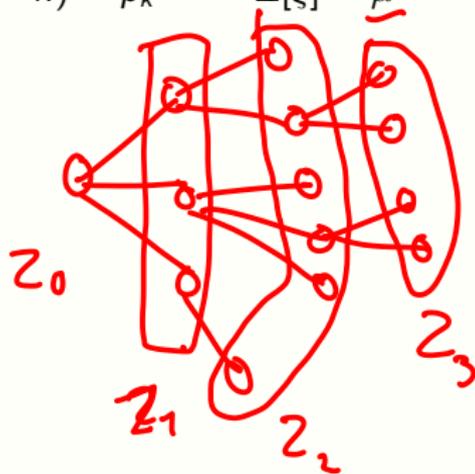
- Galton-Watson branching process
  - ▶ start with a single individual

# Proof of (1) – preliminaries

- Galton-Watson branching process

- ▶ start with a single individual
- ▶ each individual generates a number of children according to a non-negative random variable  $\xi$  with distribution  $p_k$

$$\mathbb{P}(\xi = k) = p_k \quad \mathbb{E}[\xi] = \mu \quad \text{var}(\xi) \neq 0$$



## Proof of (1) – preliminaries

- Galton-Watson branching process
  - ▶ start with a single individual
  - ▶ each individual generates a number of children according to a non-negative random variable  $\xi$  with distribution  $p_k$   
$$\mathbb{P}(\xi = k) = p_k \quad \mathbb{E}[\xi] = \mu \quad \text{var}(\xi) \neq 0$$
- Let  $Z_k$  be the number of individuals in the  $k^{\text{th}}$  generation we have  $Z_0 = 1$ ,  $Z_1 = \xi$ ,  $Z_2 = \sum_{i=1}^{Z_1} \xi^{(i)}$

## Proof of (1) – preliminaries

- Galton-Watson branching process
  - ▶ start with a single individual
  - ▶ each individual generates a number of children according to a non-negative random variable  $\xi$  with distribution  $p_k$

$$\mathbb{P}(\xi = k) = p_k \quad \mathbb{E}[\xi] = \mu \quad \text{var}(\xi) \neq 0$$

- Let  $Z_k$  be the number of individuals in the  $k^{\text{th}}$  generation  
we have  $Z_0 = 1$ ,  $Z_1 = \xi$ ,  $Z_2 = \sum_{i=1}^{Z_1} \xi^{(i)}$
- and consequently
  - ▶  $\mathbb{E}[Z_1] = \mu$

## Proof of (1) – preliminaries

- Galton-Watson branching process
  - ▶ start with a single individual
  - ▶ each individual generates a number of children according to a non-negative random variable  $\xi$  with distribution  $p_k$   
$$\mathbb{P}(\xi = k) = p_k \quad \mathbb{E}[\xi] = \mu \quad \text{var}(\xi) \neq 0$$
- Let  $Z_k$  be the number of individuals in the  $k^{\text{th}}$  generation we have  $Z_0 = 1$ ,  $Z_1 = \xi$ ,  $Z_2 = \sum_{i=1}^{Z_1} \xi^{(i)}$
- and consequently
  - ▶  $\mathbb{E}[Z_1] = \mu$
  - ▶  $\mathbb{E}[Z_2] = \mathbb{E}[\mathbb{E}[Z_2|Z_1]] = \mathbb{E}[\mu Z_1] = \underline{\mu^2}$

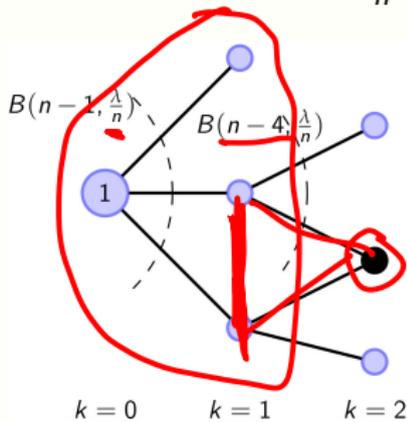
## Proof of (1) – preliminaries

- Galton-Watson branching process
  - ▶ start with a single individual
  - ▶ each individual generates a number of children according to a non-negative random variable  $\xi$  with distribution  $p_k$   
 $\mathbb{P}(\xi = k) = p_k \quad \mathbb{E}[\xi] = \mu \quad \text{var}(\xi) \neq 0$
- Let  $Z_k$  be the number of individuals in the  $k^{\text{th}}$  generation we have  $Z_0 = 1$ ,  $Z_1 = \xi$ ,  $Z_2 = \sum_{i=1}^{Z_1} \xi^{(i)}$
- and consequently
  - ▶  $\mathbb{E}[Z_1] = \mu$
  - ▶  $\mathbb{E}[Z_2] = \mathbb{E}[\mathbb{E}[Z_2|Z_1]] = \mathbb{E}[\mu Z_1] = \mu^2$
  - ▶ and by recursion, for  $k \geq 1$ , we obtain  
 $\mathbb{E}[Z_k] = \mathbb{E}[\mathbb{E}[Z_k|Z_{k-1}]] = \mathbb{E}[\mu Z_{k-1}] = \mu \cdot \mu^{k-1} = \mu^k$

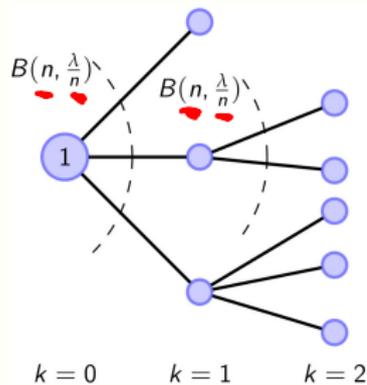
# "Proof" of (1)

$$p(m) = \frac{\lambda^m}{m!}$$

- Let  $B(n, \frac{\lambda}{n})$  denote the binomial random variable with  $n$  trials and success probability  $\frac{\lambda}{n}$



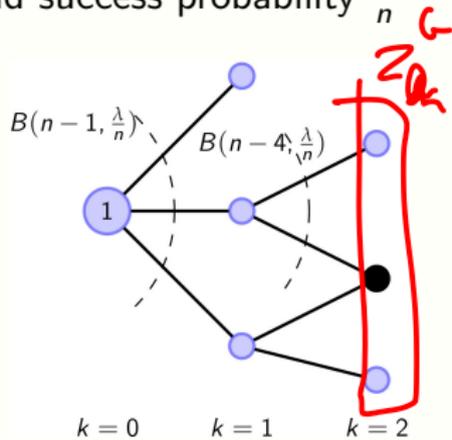
(a) ER graph process



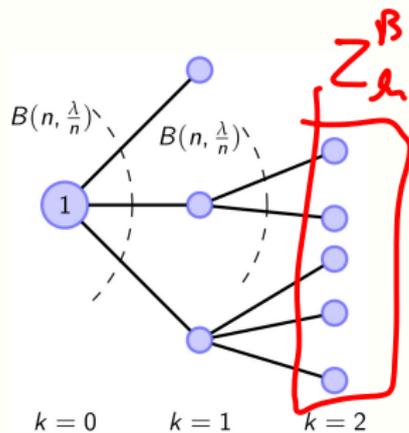
(b) branching process approx.

# "Proof" of (1)

- Let  $B(n, \frac{\lambda}{n})$  denote the binomial random variable with  $n$  trials and success probability  $\frac{\lambda}{n}$



(a) ER graph process

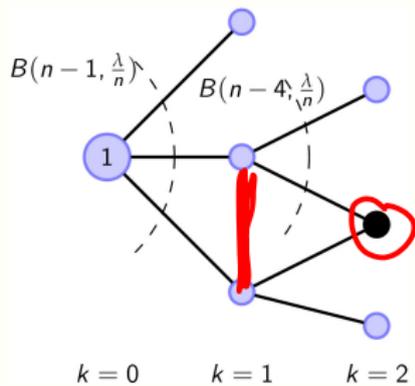


(b) branching process approx.

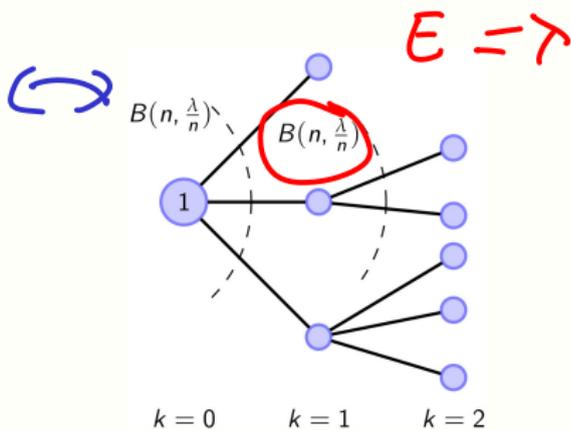
- $Z_k^G$  and  $Z_k^B$  the number of individuals in generation  $k$  for the graph process and the branching process approximation

## "Proof" of (1)

- Let  $B(n, \frac{\lambda}{n})$  denote the binomial random variable with  $n$  trials and success probability  $\frac{\lambda}{n}$



(a) ER graph process



(b) branching process approx.

- $Z_k^G$  and  $Z_k^B$  the number of individuals in generation  $k$  for the graph process and the branching process approximation
- we have  $Z_k^G \leq Z_k^B$ , for all  $k$

## "Proof" of (1)

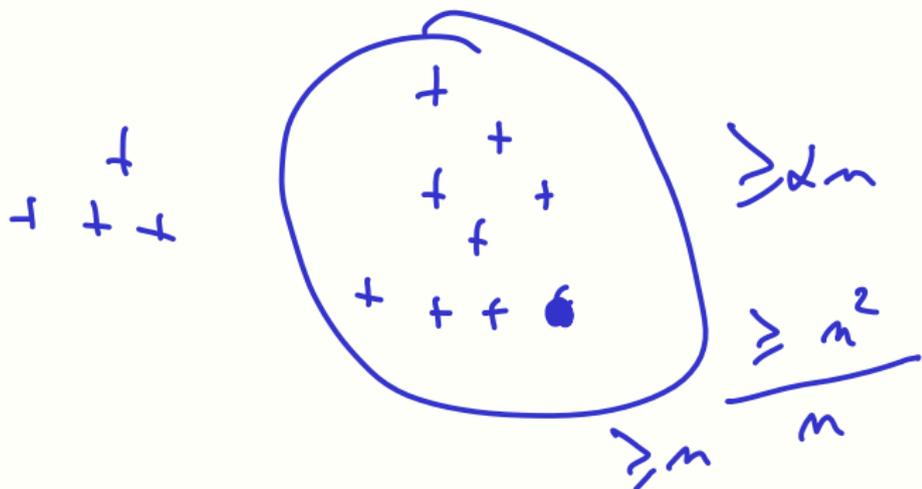
- fix  $\lambda < 1$

## "Proof" of (1)

- fix  $\lambda < 1$
- Let  $S_i$  be the number of nodes in the connected component of vertex  $i$

# "Proof" of (1)

- fix  $\lambda < 1$
- Let  $S_i$  be the number of nodes in the connected component of vertex  $i$
- we have  $\mathbb{E}[S_i] = \sum_{k \geq 0} \mathbb{E}[Z_k^G] \leq \sum_{k \geq 0} \mathbb{E}[Z_k^B] = \sum_{k \geq 0} \lambda^k = \frac{1}{1-\lambda}$



## "Proof" of (1)

- fix  $\lambda < 1$
- Let  $S_i$  be the number of nodes in the connected component of vertex  $i$
- we have  $\mathbb{E}[S_i] = \sum_k \mathbb{E}[Z_k^G] \leq \sum_k \mathbb{E}[Z_k^B] = \sum_k \lambda^k = \frac{1}{1-\lambda}$
- so if  $\lambda < 1$ , the expected size of the components of vertex  $i$  is constant  $\implies$  no giant component

## "Proof" of (1)

$$\lambda < 1 \quad \begin{array}{c} \leftarrow \quad \rightarrow \\ \lambda \\ \hline n \end{array} \quad \lambda > 1$$

- fix  $\lambda < 1$
- Let  $S_i$  be the number of nodes in the connected component of vertex  $i$
- we have  $\mathbb{E}[S_i] = \sum_k \mathbb{E}[Z_k^G] \leq \sum_k \mathbb{E}[Z_k^B] = \sum_k \lambda^k = \frac{1}{1-\lambda}$
- so if  $\lambda < 1$ , the expected size of the components of vertex  $i$  is constant  $\implies$  no giant component
- one can show (not shown here) that the size of the bigger component does not exceed  $\log n$  :

$$\forall a > 0, \mathbb{P}(\underbrace{\max_{1 \leq i \leq n} |S_i|}_{\geq a \log n}) \rightarrow 0 \text{ as } n \rightarrow +\infty$$

## Proof of (2)

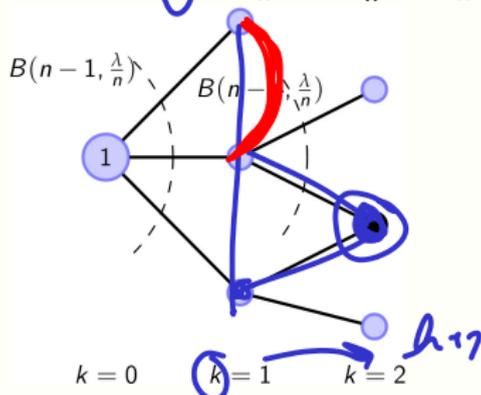
- fix  $\lambda > 1$

## Proof of (2)

- fix  $\lambda > 1$
- We want to compute  $\mathbb{E}[S_i]$  and show that it is large  
 $\implies$  we can no longer ignore conflicts

## Proof of (2)

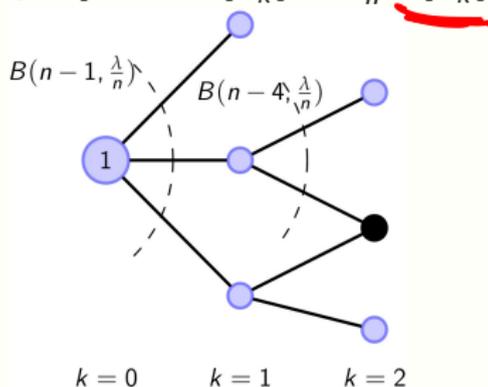
- fix  $\lambda > 1$
- We want to compute  $\mathbb{E}[S_i]$  and show that it is large  
 $\implies$  we can no longer ignore conflicts
- We claim that  $Z_k^G \approx Z_k^B$  as long as  $\lambda^k \leq cte \cdot \sqrt{n}$ 
  - ▶  $\mathbb{E}[\#\text{conflicts at stage } k] \leq n\lambda^k \mathbb{E}[Z_k^2] = n \frac{\lambda^2}{n^2} \mathbb{E}[Z_k^2]$



- ▶ what about edges between nodes of  $Z_k$ ?

## Proof of (2)

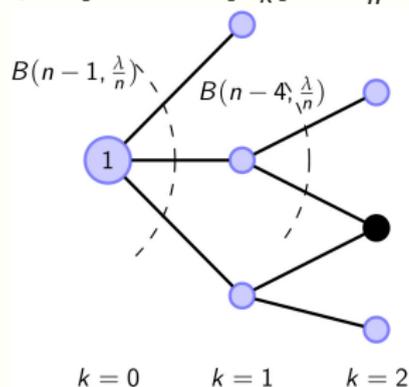
- fix  $\lambda > 1$
- We want to compute  $\mathbb{E}[S_i]$  and show that it is large  
 $\implies$  we can no longer ignore conflicts
- We claim that  $Z_k^G \approx Z_k^B$  as long as  $\lambda^k \leq cte \cdot \sqrt{n}$ 
  - ▶  $\mathbb{E}[\#\text{conflicts at stage } k] \leq np^2 \mathbb{E}[Z_k^2] = n \frac{\lambda^2}{n^2} \mathbb{E}[Z_k^2]$



- ▶ what about edges between nodes of  $Z_k$ ?
- ▶ we assume that as long as conflicts are negligible,  $Z_k$  is a Poisson variable, that is  $\text{var}(Z_k) = \lambda^k$

## Proof of (2)

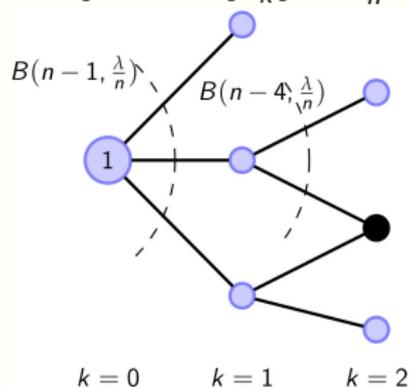
- fix  $\lambda > 1$
- We want to compute  $\mathbb{E}[S_i]$  and show that it is large  
 $\implies$  we can no longer ignore conflicts
- We claim that  $Z_k^G \approx Z_k^B$  as long as  $\lambda^k \leq cte \cdot \sqrt{n}$ 
  - ▶  $\mathbb{E}[\#\text{conflicts at stage } k] \leq np^2 \mathbb{E}[Z_k^2] = n \frac{\lambda^2}{n^2} \mathbb{E}[Z_k^2]$



- ▶ what about edges between nodes of  $Z_k$ ?
- ▶ we assume that as long as conflicts are negligible,  $Z_k$  is a Poisson variable, that is  $\text{var}(Z_k) = \lambda^k$
- ▶ we obtain  $\mathbb{E}[Z_k^2] = \text{var}(Z_k) + \mathbb{E}[Z_k]^2 = \lambda^k + \lambda^{2k} \sim \lambda^{2k}$

## Proof of (2)

- fix  $\lambda > 1$
- We want to compute  $\mathbb{E}[S_i]$  and show that it is large  
 $\implies$  we can no longer ignore conflicts
- We claim that  $Z_k^G \approx Z_k^B$  as long as  $\lambda^k \leq cte \cdot \sqrt{n}$ 
  - ▶  $\mathbb{E}[\#\text{conflicts at stage } k] \leq np^2 \mathbb{E}[Z_k^2] = n \frac{\lambda^2}{n^2} \mathbb{E}[Z_k^2]$



- ▶ what about edges between nodes of  $Z_k$ ?
  - ▶ we assume that as long as conflicts are negligible,  $Z_k$  is a Poisson variable, that is  $var(Z_k) = \lambda^k$
  - ▶ we obtain  $\mathbb{E}[Z_k^2] = var(Z_k) + \mathbb{E}[Z_k]^2 = \lambda^k + \lambda^{2k} \sim \lambda^{2k}$
- $\implies \mathbb{E}[\#\text{conflicts}]$  becomes  $\Omega(1)$  only when  $\lambda^k \approx \sqrt{n}$

## Proof of (2)

$$\begin{aligned} \bullet \mathbb{E}[S_i] &= \sum_k \mathbb{E}[Z_k^G] \geq \sum_{k \leq \log_\lambda(\sqrt{n})} \mathbb{E}[Z_k^B] = \sum_{k \leq \log_\lambda(\sqrt{n})} \lambda^k \\ &\geq \frac{1 - \lambda^{\log_\lambda(\sqrt{n})}}{1 - \lambda} \geq \sqrt{n} \end{aligned}$$

## Proof of (2)

- $\mathbb{E}[S_i] = \sum_k \mathbb{E}[Z_k^G] \geq \sum_{k \leq \log_\lambda(\sqrt{n})} \mathbb{E}[Z_k^B] = \sum_{k \leq \log_\lambda(\sqrt{n})} \lambda^k$   
 $\geq \frac{1 - \lambda^{\log_\lambda(\sqrt{n})}}{1 - \lambda} \geq \sqrt{n}$
- Let us assume again that  $|Z_k|$  follows a Poisson law of parameter  $\lambda^k$ 
  - ▶ we then have  $\mathbb{P}(|Z_k| - \lambda^k \geq x) \leq 2e^{-\frac{x^2}{2(\lambda^k + x)}}$
  - ▶ which gives for  $x = \sqrt{\lambda^k}$ ,  $\mathbb{P}(|Z_k| - \lambda^k \geq \sqrt{\lambda^k}) \leq 2e^{-\frac{1}{3}}$

## Proof of (2)

$$\begin{aligned} \bullet \mathbb{E}[S_i] &= \sum_k \mathbb{E}[Z_k^G] \geq \sum_{k \leq \log_\lambda(\sqrt{n})} \mathbb{E}[Z_k^B] = \sum_{k \leq \log_\lambda(\sqrt{n})} \lambda^k \\ &\geq \frac{1 - \lambda^{\log_\lambda(\sqrt{n})}}{1 - \lambda} \geq \sqrt{n} \end{aligned}$$

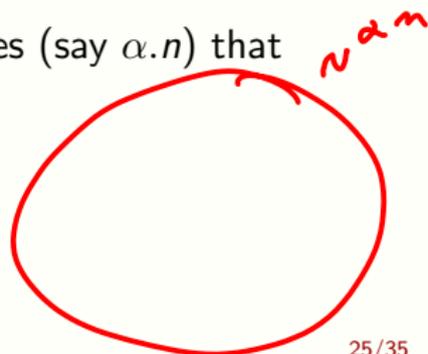
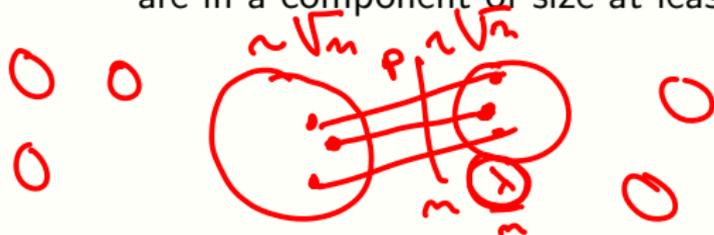
- Let us assume again that  $|Z_k|$  follows a Poisson law of parameter  $\lambda^k$

▶ we then have  $\mathbb{P}(|Z_k| - \lambda^k \geq x) \leq 2e^{-\frac{x^2}{2(\lambda^k + x)}}$

▶ which gives for  $x = \sqrt{\lambda^k}$ ,  $\mathbb{P}(|Z_k| - \lambda^k \geq \sqrt{\lambda^k}) \leq 2e^{-\frac{1}{3}}$

- for large  $n$ , we obtain  $\mathbb{P}(|S_i| \geq \frac{\sqrt{n}}{2}) \geq cte$

$\implies$  there is a constant fraction of the nodes (say  $\alpha \cdot n$ ) that are in a component of size at least  $\frac{\sqrt{n}}{2}$



## Proof of (2)

- Assume there is more than one component of size  $\frac{\sqrt{n}}{2}$ , we will show that the probability this happens  $\rightarrow 0$  when  $n \rightarrow +\infty$

## Proof of (2)

- Assume there is more than one component of size  $\frac{\sqrt{n}}{2}$ , we will show that the probability this happens  $\rightarrow 0$  when  $n \rightarrow +\infty$
- Let  $C_1$  be the smallest of these components and let  $A$  be the union of all these components

## Proof of (2)

- Assume there is more than one component of size  $\frac{\sqrt{n}}{2}$ , we will show that the probability this happens  $\rightarrow 0$  when  $n \rightarrow +\infty$
- Let  $C_1$  be the smallest of these components and let  $A$  be the union of all these components
  - ▶ we denote  $|C_1| = k \geq \frac{\sqrt{n}}{2}$

## Proof of (2)

- Assume there is more than one component of size  $\frac{\sqrt{n}}{2}$ , we will show that the probability this happens  $\rightarrow 0$  when  $n \rightarrow +\infty$
- Let  $C_1$  be the smallest of these components and let  $A$  be the union of all these components
  - ▶ we denote  $|C_1| = k \geq \frac{\sqrt{n}}{2}$
  - ▶  $\mathbb{P}(C_1 \text{ not connected to } A \setminus C_1) = (1 - p)^{k(|A| - k)} \leq (1 - \frac{\lambda}{n})^{\frac{\alpha n \sqrt{n}}{4}} \leq e^{-\frac{\lambda}{n} \cdot \frac{\alpha n \sqrt{n}}{4}} = e^{-\frac{\lambda \alpha \sqrt{n}}{4}} \rightarrow 0$  when  $n \rightarrow +\infty$

## Proof of (2)

- Assume there is more than one component of size  $\frac{\sqrt{n}}{2}$ , we will show that the probability this happens  $\rightarrow 0$  when  $n \rightarrow +\infty$
- Let  $C_1$  be the smallest of these components and let  $A$  be the union of all these components
  - ▶ we denote  $|C_1| = k \geq \frac{\sqrt{n}}{2}$
  - ▶  $\mathbb{P}(C_1 \text{ not connected to } A \setminus C_1) = (1 - p)^{k(|A| - k)} \leq (1 - \frac{\lambda}{n})^{\frac{\alpha n \sqrt{n}}{4}} \leq e^{-\frac{\lambda}{n} \cdot \frac{\alpha n \sqrt{n}}{4}} = e^{-\frac{\lambda \alpha \sqrt{n}}{4}} \rightarrow 0$  when  $n \rightarrow +\infty$
- this means that the probability that the vertices of  $A$  are grouped in a single connected component  $\rightarrow 1$  when  $n \rightarrow +\infty$

## Proof of (2)

- Assume there is more than one component of size  $\frac{\sqrt{n}}{2}$ , we will show that the probability this happens  $\rightarrow 0$  when  $n \rightarrow +\infty$
- Let  $C_1$  be the smallest of these components and let  $A$  be the union of all these components
  - ▶ we denote  $|C_1| = k \geq \frac{\sqrt{n}}{2}$
  - ▶  $\mathbb{P}(C_1 \text{ not connected to } A \setminus C_1) = (1 - p)^{k(|A| - k)} \leq (1 - \frac{\lambda}{n})^{\frac{\alpha n \sqrt{n}}{4}} \leq e^{-\frac{\lambda}{n} \cdot \frac{\alpha n \sqrt{n}}{4}} = e^{-\frac{\lambda \alpha \sqrt{n}}{4}} \rightarrow 0$  when  $n \rightarrow +\infty$
- this means that the probability that the vertices of  $A$  are grouped in a single connected component  $\rightarrow 1$  when  $n \rightarrow +\infty$
- since  $|A| \geq \alpha \cdot n$ , this constitutes a giant component

## Size of the giant component

- Let  $G = G_{n-1,p}$  be an ER graph with  $p(n) = \frac{\lambda}{n}$  with  $\lambda > 1$

## Size of the giant component

- Let  $G = G_{n-1,p}$  be an ER graph with  $p(n) = \frac{\lambda}{n}$  with  $\lambda > 1$
- Add a  $n^{\text{th}}$  vertex to  $G$  and connect it to the rest of the vertices with probability  $p(n)$  and denote  $G'$  the resulting graph

## Size of the giant component

- Let  $G = G_{n-1, p}$  be an ER graph with  $p(n) = \frac{\lambda}{n}$  with  $\lambda > 1$
- Add a  $n^{\text{th}}$  vertex to  $G$  and connect it to the rest of the vertices with probability  $p(n)$  and denote  $G'$  the resulting graph
- We denote  $\rho$  the fraction of vertices that are not in the giant component and we assume that, for large  $n$ ,  $\rho$  is the same in  $G$  and  $G'$

## Size of the giant component

- Let  $G = G_{n-1, \rho}$  be an ER graph with  $\rho(n) = \frac{\lambda}{n}$  with  $\lambda > 1$
- Add a  $n^{\text{th}}$  vertex to  $G$  and connect it to the rest of the vertices with probability  $\rho(n)$  and denote  $G'$  the resulting graph
- We denote  $\rho$  the fraction of vertices that are not in the giant component and we assume that, for large  $n$ ,  $\rho$  is the same in  $G$  and  $G'$
- vertex  $n$  is not in the giant component iff none of its neighbours are
  - ▶ This gives  $\rho = \sum_{d \geq 0} P_d \rho^d = \Phi(\rho)$

## Size of the giant component

- Let  $G = G_{n-1, \rho}$  be an ER graph with  $\rho(n) = \frac{\lambda}{n}$  with  $\lambda > 1$
- Add a  $n^{\text{th}}$  vertex to  $G$  and connect it to the rest of the vertices with probability  $\rho(n)$  and denote  $G'$  the resulting graph
- We denote  $\rho$  the fraction of vertices that are not in the giant component and we assume that, for large  $n$ ,  $\rho$  is the same in  $G$  and  $G'$
- vertex  $n$  is not in the giant component iff none of its neighbours are
  - ▶ This gives  $\rho = \sum_{d \geq 0} P_d \rho^d = \Phi(\rho)$
- The analysis of function  $\Phi$  shows that it has a unique fixed point  $\rho^* \in ]0, 1[$

## Mean distance at the connectivity threshold

(very) roughly speaking

- $\mathbb{E}[\frac{\#\text{conflicts at stage } k}{n}] \leq np^2 \mathbb{E}[Z_k^2] \approx \frac{\log^2 n}{n} \cdot \log^{2k} n = \frac{\log^{2(k+1)} n}{n}$

## Mean distance at the connectivity threshold

(very) roughly speaking

- $\mathbb{E}[\frac{\#\text{conflicts at stage } k}{n}] \leq np^2 \mathbb{E}[Z_k^2] \approx \frac{\log^2 n}{n} \cdot \log^{2k} n =$ 
  - ▶ What about edges between vertices of  $Z_k$ ?

# Mean distance at the connectivity threshold

(very) roughly speaking

- $\mathbb{E}[\frac{\#\text{conflicts at stage } k}{n}] \leq np^2 \mathbb{E}[Z_k^2] \approx \frac{\log^2 n}{n} \cdot \log^{2k} n = \frac{\log^{2(k+1)} n}{n}$ 
  - ▶ What about edges between vertices of  $Z_k$ ?
- Conflicts are negligible until  $\frac{\log^{2(k+1)} n}{n} = 1$ , that is  $k = \frac{\log n}{2 \log \log n} - 1$

# Mean distance at the connectivity threshold

(very) roughly speaking

- $\mathbb{E}[\frac{\#\text{conflicts at stage } k}{n}] \leq np^2 \mathbb{E}[Z_k^2] \approx \frac{\log^2 n}{n} \cdot \log^{2k} n = \frac{\log^{2(k+1)} n}{n}$ 
  - ▶ What about edges between vertices of  $Z_k$ ?
- Conflicts are negligible until  $\frac{\log^{2(k+1)} n}{n} = 1$ , that is  $k = \frac{\log n}{2 \log \log n} - 1$
- Then  $|S_i| \approx (\log n)^{\frac{\log n}{2 \log \log n} - 1} = \frac{\sqrt{(\log n)^{\frac{\log n}{\log \log n}}}}{\log n} = \frac{\sqrt{n}}{\log n}$

## Mean distance at the connectivity threshold

(very) roughly speaking

- $\mathbb{E}[\frac{\#\text{conflicts at stage } k}{n}] \leq np^2 \mathbb{E}[Z_k^2] \approx \frac{\log^2 n}{n} \cdot \log^{2k} n = \frac{\log^{2(k+1)} n}{n}$ 
  - ▶ What about edges between vertices of  $Z_k$ ?
- Conflicts are negligible until  $\frac{\log^{2(k+1)} n}{n} = 1$ , that is  $k = \frac{\log n}{2 \log \log n} - 1$
- Then  $|S_i| \approx (\log n)^{\frac{\log n}{2 \log \log n} - 1} = \frac{\sqrt{(\log n)^{\frac{\log n}{\log \log n}}}}{\log n} = \frac{\sqrt{n}}{\log n}$
- One can cover the vertex set by approx.  $\sqrt{n} \log n$  balls of size approx.  $\frac{\sqrt{n}}{\log n}$

# Mean distance at the connectivity threshold

(very) roughly speaking

- $\mathbb{E}[\frac{\#\text{conflicts at stage } k}{n}] \leq np^2 \mathbb{E}[Z_k^2] \approx \frac{\log^2 n}{n} \cdot \log^{2k} n = \frac{\log^{2(k+1)} n}{n}$ 
  - ▶ What about edges between vertices of  $Z_k$ ?
- Conflicts are negligible until  $\frac{\log^{2(k+1)} n}{n} = 1$ , that is  $k = \frac{\log n}{2 \log \log n} - 1$
- Then  $|S_i| \approx (\log n)^{\frac{\log n}{2 \log \log n} - 1} = \frac{\sqrt{(\log n)^{\frac{\log n}{\log \log n}}}}{\log n} = \frac{\sqrt{n}}{\log n}$
- One can cover the vertex set by approx.  $\sqrt{n} \log n$  balls of size approx.  $\frac{\sqrt{n}}{\log n}$
- the probability for two such balls not to be connected by an edge is  $(1 - p)^{\frac{n}{\log^2 n}} \leq e^{-\frac{1}{\log n}}$

## Mean distance at the connectivity threshold

(very) roughly speaking

- $\mathbb{E}[\frac{\#\text{conflicts at stage } k}{n}] \leq np^2 \mathbb{E}[Z_k^2] \approx \frac{\log^2 n}{n} \cdot \log^{2k} n = \frac{\log^{2(k+1)} n}{n}$ 
  - ▶ What about edges between vertices of  $Z_k$ ?
- Conflicts are negligible until  $\frac{\log^{2(k+1)} n}{n} = 1$ , that is  $k = \frac{\log n}{2 \log \log n} - 1$
- Then  $|S_i| \approx (\log n)^{\frac{\log n}{2 \log \log n} - 1} = \frac{\sqrt{(\log n)^{\frac{\log n}{\log \log n}}}}{\log n} = \frac{\sqrt{n}}{\log n}$
- One can cover the vertex set by approx.  $\sqrt{n} \log n$  balls of size approx.  $\frac{\sqrt{n}}{\log n}$
- the probability for two such balls not to be connected by an edge is  $(1 - p)^{\frac{n}{\log^2 n}} \leq e^{-\frac{1}{\log n}}$
- so the proba for them to be connected is at least  $1 - e^{-\frac{1}{\log n}} \sim \frac{1}{\log n}$

## Mean distance at the connectivity threshold

- let us write  $N = \sqrt{n} \log n$ , and call  $\tilde{G}$  the graph on the  $N$  balls that cover  $G$

## Mean distance at the connectivity threshold

- let us write  $N = \sqrt{n} \log n$ , and call  $\tilde{G}$  the graph on the  $N$  balls that cover  $G$
- we have  $\log N \sim \frac{\log n}{2}$  and  $\tilde{G}$  contains an ER graph on  $N$  vertices with  $\tilde{p} = \frac{1}{2 \log N}$

## Mean distance at the connectivity threshold

- let us write  $N = \sqrt{n} \log n$ , and call  $\tilde{G}$  the graph on the  $N$  balls that cover  $G$
- we have  $\log N \sim \frac{\log n}{2}$  and  $\tilde{G}$  contains an ER graph on  $N$  vertices with  $\tilde{p} = \frac{1}{2 \log N}$
- In  $\tilde{G}$  the probability for two given nodes to be at distance more than 2 is at most  $(1 - \frac{1}{2 \log N})^{N-1} \leq e^{-\frac{N-1}{2 \log N}} \rightarrow 0$  when  $N \rightarrow +\infty$ .

## Mean distance at the connectivity threshold

- let us write  $N = \sqrt{n} \log n$ , and call  $\tilde{G}$  the graph on the  $N$  balls that cover  $G$
- we have  $\log N \sim \frac{\log n}{2}$  and  $\tilde{G}$  contains an ER graph on  $N$  vertices with  $\tilde{p} = \frac{1}{2 \log N}$
- In  $\tilde{G}$  the probability for two given nodes to be at distance more than 2 is at most  $(1 - \frac{1}{2 \log N})^{N-1} \leq e^{-\frac{N-1}{2 \log N}} \rightarrow 0$  when  $N \rightarrow +\infty$ .
- Therefore, between any two vertices of  $G$  there exists with probability tending to 1 when  $n \rightarrow +\infty$  a path of length  $(\frac{\log n}{2 \log \log n} - 1) + 1 + 2(\frac{\log n}{2 \log \log n} - 1) + 1 + (\frac{\log n}{2 \log \log n} - 1) \leq \frac{2 \log n}{\log \log n}$

## Configuration model – Molloy & Reed 1995

Input : an arbitrary degree distribution

Output : a random graph with this degree distribution

# Configuration model – Molloy & Reed 1995

Input : an arbitrary degree distribution

Output : a random graph with this degree distribution

Generation process :

1. Assign a fixed number of semi-links to each node (according to the input degree distribution)
2. Pair the semi-links uniformly at random
3. Remove self-loops and multiple edges

# Configuration model – Molloy & Reed 1995

Input : an arbitrary degree distribution

Output : a random graph with this degree distribution

Generation process :

1. Assign a fixed number of semi-links to each node (according to the input degree distribution)
2. Pair the semi-links uniformly at random
3. Remove self-loops and multiple edges

What degree distribution should we take as parameter ?

- The degree distribution of some real-world network
- A mathematically defined one, powerlaw  $\mathbb{P}(k) \sim k^{-\alpha}$ .

## Configuration model : implementation and complexity

- Put the semi-links in a table of size  $2m$
- Pick  $m$  times two of them uniformly at random

# Properties of the configuration model

Four properties to check :

- Low global density ✓
  - ▶ the degree distribution is the parameter of the model and controls  $m$  :  $m = \frac{\sum_{0 \leq k \leq n-1} k \cdot N_k}{2}$

# Properties of the configuration model

Four properties to check :

- Low global density ✓
- Short distances ✓

Expansion property :

- ▶ Degree of the extremity of one edge :

$$\mathbb{P}(d^\circ(\text{ext}) = k') = \frac{k' \mathbb{P}(k')}{\langle k \rangle}$$

- ▶ Probability that following one edge leads to  $k$  new vertices :

$$q(k) = \mathbb{P}(d^\circ(\text{ext}) = k + 1)$$

- ▶ Expected number of new vertices following one edge :

$$\sum_k k q(k) = \frac{\langle k^2 \rangle - \langle k \rangle}{\langle k \rangle}$$

# Properties of the configuration model

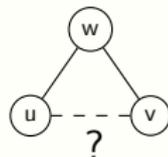
Four properties to check :

- Low global density ✓
- Short distances ✓
- Heterogeneous degrees ✓
  - ▶ the degree distribution is the parameter of the model

# Properties of the configuration model

Four properties to check :

- Low global density ✓
- Short distances ✓
- Heterogeneous degrees ✓
- High local density

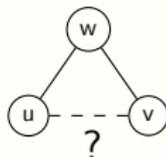


- ▶ Probability to have a link between  $u$  and  $k$  with  $d^\circ(u) = k$  and  $d^\circ(v) = k'$  :  $\mathbb{P}(uv|kk') = \frac{kk'}{\langle k \rangle N}$

# Properties of the configuration model

Four properties to check :

- Low global density ✓
- Short distances ✓
- Heterogeneous degrees ✓
- High local density



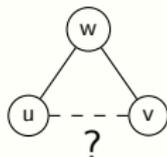
- ▶ Probability to have a link between  $u$  and  $k$  with  $d^\circ(u) = k$  and  $d^\circ(v) = k'$  :  $\mathbb{P}(uv|kk') = \frac{kk'}{\langle k \rangle N}$
- ▶ Probability to have a link between  $u$  and  $v$  :

$$\begin{aligned}\mathbb{P}(\text{triangle}) &= \frac{\sum_{k \geq 1} \sum_{k' \geq 1} \frac{kk'}{\langle k \rangle N} q(k)q(k')}{\langle k \rangle N} \\ &= \frac{1}{\langle k \rangle N} \sum_{k \geq 1} kq(k) \sum_{k' \geq 1} k'q(k') \\ &= \frac{1}{N} \frac{(\langle k^2 \rangle - \langle k \rangle^2)}{\langle k \rangle^3}\end{aligned}$$

# Properties of the configuration model

Four properties to check :

- Low global density ✓
- Short distances ✓
- Heterogeneous degrees ✓
- High local density ✗



- ▶ Probability to have a link between  $u$  and  $k$  with  $d^\circ(u) = k$  and  $d^\circ(v) = k'$  :  $\mathbb{P}(uv|kk') = \frac{kk'}{\langle k \rangle N}$
- ▶ Probability to have a link between  $u$  and  $v$  :

$$\begin{aligned}\mathbb{P}(\text{triangle}) &= \frac{1}{\langle k \rangle N} \sum_{k \geq 1} \sum_{k' \geq 1} \frac{kk'}{\langle k \rangle N} q(k)q(k') \\ &= \frac{1}{\langle k \rangle N} \sum_{k \geq 1} kq(k) \sum_{k' \geq 1} k'q(k') \\ &= \frac{1}{N} \frac{(\langle k^2 \rangle - \langle k \rangle^2)}{\langle k \rangle^3} \\ &\rightarrow 0 \text{ when } N \rightarrow +\infty\end{aligned}$$